

1 Function of Several Random Variables

Let $X_1, X_2 \dots X_n$ be a given set of random variables and suppose another random variable Z is defined via $X_1, X_2 \dots X_n$ by

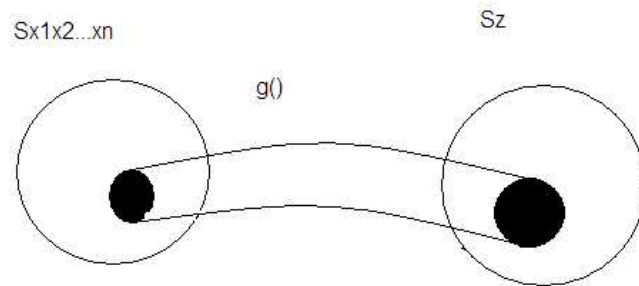
$$Z := g(X_1, X_2 \dots X_n)$$

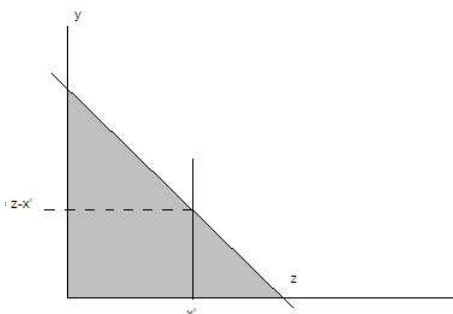
We are interested in determining the distribution of $(X_1 \dots X_n)$ and function g .

First Principle

For any $B \subseteq S_Z$,

$$P[Z \in B] = P[(X_1, X_2 \dots X_n) \in g^{-1}(B)]$$





2 Real-Valued Functions 'g'

Summation of independent random variables

Let X and Y be independent random variables (RV's) with PDF f_x and f_y respectively. Suppose $Z = X + Y$. Find f_z in terms of f_x and f_y .

We will determine f_z via F_Z :

$$\begin{aligned} F_Z(z) &= P[Z \leq z] \\ &= P[X + Y \leq z] \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{z-x'} f_{xy}(x', y') dx' \end{aligned}$$

Let $h(z, x') = \int_{-\infty}^{z-x'} f_{xy}(x', y') dx'$

$$\begin{aligned} F_Z(z) &= \int_{-\infty}^{\infty} h(z, x') f_x(x') dx' \\ f'_z(z) &= \frac{dF_Z(z)}{dz} \\ &= \frac{d}{dz} \int_{-\infty}^{\infty} h(z, x') f_x(x') dx' \end{aligned}$$

$$\begin{aligned}
&= \int_{-\infty}^{\infty} \frac{d}{dz} h(z, x') dx' \\
&= \frac{d}{dz} \int_{-\infty}^{z-x'} f_{xy}(x', y') dy' \\
&= f_{xy}(x', z - x')
\end{aligned}$$

$$\text{then } f_Z(z) = \int_{-\infty}^{\infty} f_{xy}(x', z - x') dx'$$

Note: the above equation is true whether or not X and Y are independent.

Since X and Y are independent, $f_{xy}(x, y) = f_x(x)f_y(y)$

$$f_Z(z) = \int_{-\infty}^{\infty} f_x(x')f_y(z - x')dx' = f_x(z) * f_y(z)$$

In general, if $z = x_1 + x_2 + \dots + x_n$, where x_1, x_2, \dots, x_n are independent, then:

$$f_z(z) = f_{x_1}(z) * f_{x_2}(z) * \dots * f_{x_n}(z)$$

Another more intuitive but less rigorous method to see $f_{x+y}(z) = f_x(z) * f_y(z)$ for independent X and Y is the following:

$$f_z(z) = \int_{x+y=z} f_{xy}(x, y) dx dy = \int_{-\infty}^{\infty} f_x(x)f_y(z - x)dx = f_x(z) * f_y(z)$$

In fact, such results hold for continuous, discrete, and even mixed-type RV's. The proof using characteristic functions will show this.

$$\Phi(\omega) = E[e^{j\omega Z}] = E[e^{j\omega(X+Y)}] = E[e^{j\omega X} \bullet e^{j\omega Y}]$$

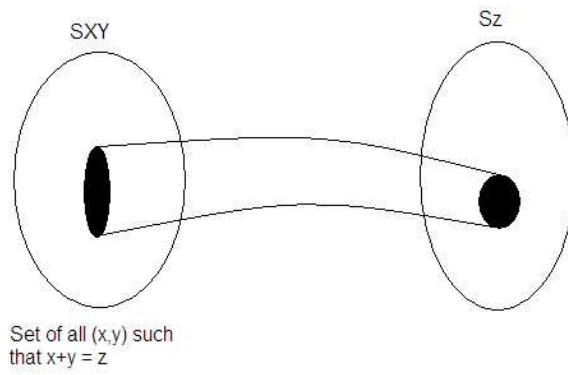
Aside: if X and Y are independent, then for any two functions g and h , $g(X)$ and $h(Y)$ are independent.

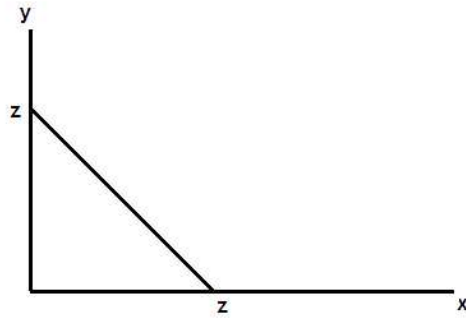
Then $e^{j\omega X}$ is independent of $e^{j\omega Y}$

Aside: if X and Y are independent, then $E[XY] = E[X]E[Y]$

$$\text{Then } \Phi_2(\omega) = E[e^{j\omega X}]E[e^{j\omega Y}] = \Phi_x(\omega) \bullet \Phi_y(\omega)$$

Thus, $f_2(y) = f_x(z) * f_y(z)$





Example

Suppose that independent messages are arriving at a station where the inter-arrival duration is modelled as an exponential distribution with parameter λ , (ie, λ is the average number of messages arriving per unit time). Suppose now is the time 0; Let Z be the time at which the fifth message will arrive. Find the PDF of Z .

Solution

(I) Let x_i be the time interval between message i and message $i - 1$. (Therefore, x_1 is the time of the first message.)

Then $z = x_1 + x_2 + \dots + x_5$

$$f_z(z) = f_{x_1}(z) * \dots * f_{x_5}(z)$$

Since x_1, x_2, \dots, x_5 are independent, we will use \hat{f} to denote the fourier transform of f .

$$\hat{f}(\omega) = \hat{f}_{x_1}(\omega) \bullet \dots \bullet \hat{f}_{x_5}(\omega)$$

Since

$$\hat{f}_{x_1}(\omega) = \hat{f}_{x_2}(\omega) = \dots = \hat{f}_{x_5}(\omega) = F[\lambda e^{\lambda t}]$$

Aside

$$e^{-\lambda t} u(t) \xleftrightarrow{F} \frac{1}{(\lambda + j\omega)}$$

$$\frac{t^{(n-1)}}{(n-1)!} \xleftrightarrow{F} \frac{1}{(\lambda + j\omega)^n}$$

Then

$$\hat{f}_Z(\omega) = \left(\frac{\lambda}{\lambda + j\omega}\right)^5 = \frac{\lambda^5}{(\lambda + j\omega)^5}$$

Inverse Fourier Transform

$$f_Z(\omega) = \frac{\lambda^5 z^4}{4!} e^{-\lambda z} u(z)$$

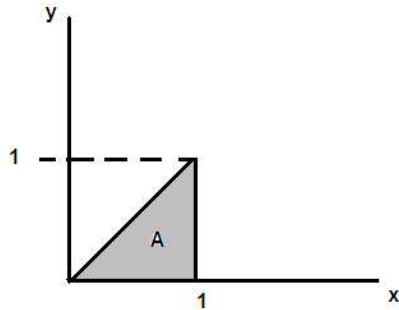
This is called m-Erlang distribution with m=5

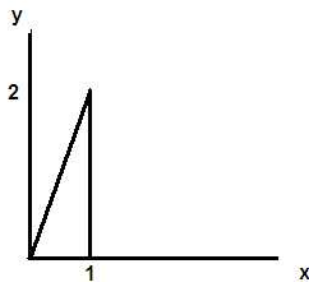
(II) Finding PDF of function of several random variable's via conditional PDF.

Example

Let (x,y) be uniformly distributed over region A.

Let $Z := XY$, Find PDF of Z



**Solution**

$$f_X(X) = \begin{cases} 2x, & x \in [0, 1] \\ 0, & \text{otherwise} \end{cases}$$

We will use the fact that

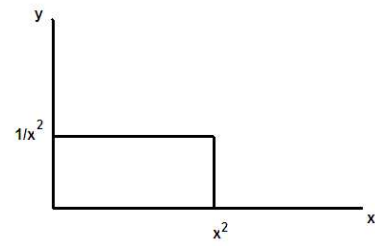
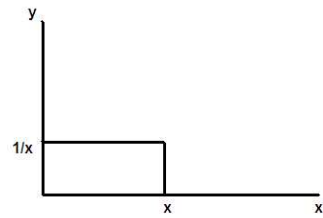
$$f_Z(z) = \int_{-\infty}^{\infty} f_{Z|X=x}(z|x) \cdot f_X(x) dx$$

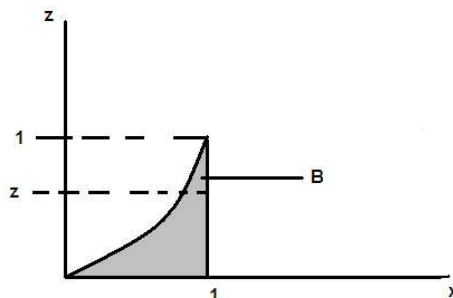
$$f_{Z|X}(z|x) = f_{Z|X=x}(z)$$

note given $X = x$, then $Z = xY$, where Y follows the distribution $f_{Y|X=x}(y)$

$$f_{Y|X=x}(y) = \begin{cases} \frac{1}{x} & x \in [0, x] \\ 0, & \text{otherwise} \end{cases}$$

As Z is now simply a scaled version of Y the distribution of Z (conditioned on $X = x$)





$$f_{Z|X=x}(z) = \begin{cases} \frac{1}{x^2} & z \in [0, x^2] \\ 0, & \text{otherwise} \end{cases}$$

$$f_{zx}(z, x) = f_{z|x}(z|x)f_x(x)$$

$$f_{zx}(z, x) = \begin{cases} \frac{2x}{x^2} = \frac{2}{x}, & x \in [0, 1], z \in [0, x^2] \\ 0, & \text{otherwise} \end{cases}$$

$$\begin{aligned} f_z(z) &= \int_{-\infty}^{\infty} f_{zx}(z, x) dx \\ &= \int_{\sqrt{z}}^1 \frac{2}{x} dx \\ &= 2\ln(x) \Big|_{\sqrt{z}}^1 \\ &= -2\ln\sqrt{z} \\ &= -\ln z \text{ for } z \in [0, 1] \end{aligned}$$